

# IMPACT OF TRANSFER PRICING ON FOREIGN DIRECT INVESTMENT IN ROMANIA

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**Abstract** *The purpose of the study presented in this paper is to analyse the impact of transfer pricing on foreign direct investment (FDI) in Romania. For attaining this goal, we performed a simple linear regression by the least squares method to study the impact of adjustments of tax obligations in the field of transfer pricing on foreign direct investment in the period 2011-2019. We have proved, from a statistical point of view, that there is a relationship between foreign direct investment and adjustments to tax liabilities resulting from transfer pricing.*

**Key words:** *transfer prices, foreign direct investment, multinational companies, simple linear regression equation*

## 1. Introduction

Transfer pricing is interdisciplinary in nature, being under the influence of accounting, taxation, economics and law.

Studies in this field show the economists' constant concern over analysing how multinational companies move their profits, the link between tax rates and reported profits, and the effects of local regulations on decisions to locate multinational companies. We intend to use an econometric model to test the relationship between the level of adjustments to tax liabilities related to transfer prices and foreign direct investment in Romania.

The conclusions reached in this paper are based on the documentation on the existing bibliography in the field, as well as on the use of statistical-mathematical and econometric methods, more exactly a simple linear regression - the least squares method, built on econometric software EViews 7.2. and Excel statistical-mathematical software.

The paper is organized as follows: the following section assesses the link between FDI and multinational companies as main generators of transfer pricing. Section 3 develops our econometric model for testing the relationship between the level of adjustments of

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fiscal obligations related to transfer pricing and FDI in Romania, and the last section offers the conclusions.

## 2. Transfer Pricing and Relationship between Foreign Direct Investment and Multinational Companies

We consider first the links between FDI and multinational companies (MNCs), sometimes called multinational enterprises (MNEs), as the main generators of intra-group transactions and implicitly transfer prices. There are strong links between foreign direct investment and multinational companies, the former being the main financial source for setting up multinational companies.

According to the definition given by the NBR, direct investments represent long-term investment relations between resident and non-resident entities, which involve the exercise by investors of significant managerial influence in the enterprises in which they have invested.

Taking into account the fact that the FDI balance represents the value of all foreign direct investments that have been accumulated until a certain moment (in the analysed case - at the end of the year), the balance of foreign direct investments in Romania (for the period 2011-2019) is as follows:

Table 1  
*Evolution of FDI balance in GDP for the period 2011-2019 in Romania (mil EUR)*

Components	2011	2012	2013	2014	2015	2016	2017	2018	2019
FDI Balance Total *	53.723	57.851	60.639	61.999	64.663	70.742	75.851	81.124	88.304
Equity	37.081	39.393	40.700	43.243	45.098	48.964	52.746	57.479	61.352
Debt instruments **	16.642	18.458	19.939	18.756	19.565	21.778	23.105	23.645	26.952
GDP*	131.963	133.246	143.802	150.428	160.328	170.378	187.801	204.684	223.342
FDI/GDP (%) Balance	40,7	43,4	42,2	41,2	40,3	41,5	40,4	39,6	39,5

Source: *National Bank of Romania, Foreign direct investment in Romania in 2019 (BNR, 2019)*

*Notes:*

\* - FDI balances for the period 2011-2012 were recalculated according to the methodology provided by the IMF Manual Balance of Payments and International Investment Position (BPM6)

- FDI balances for the period 2013 - 2016 were revised in December 2019, within the benchmark revision (more information is available at <http://www.bnr.ro/Procesul-de-revizuire-a-datelor-statistice-20794.aspx>)

\*\* debts minus receivables (claims) in relation to foreign direct investors and companies in their group

FDI flows, representing new investments that entered the economy in one year, evolved in Romania in the period 2011-2019 as follows:

Table 2  
*Evolution of FDI flow in GDP for the period 2011-2019 in Romania (million EUR)*

Components	2011	2012	2013	2014	2015	2016	2017	2018	2019
Total*	1.700	2.489	2.712	2.421	3.461	4.517	4.797	5.266	5.173
Contribution to equity	4.002	2.676	2.765	4.222	3.085	3.202	2.235	2.973	2.238
Reinvested profit	-2.497	-1.846	-337	-1.376	510	1.138	1.733	2.573	2.783
Debt instruments	195	1.659	285	-425	-134	176	829	-280	152
GDP*	131.963	133.246	143.802	150.428	160.328	170.378	187.801	204.684	223.342
FDI/GDP (%) flow	1,3	1,9	1,9	1,6	2,2	2,7	2,6	2,6	2,3

Source: *National Bank of Romania, Foreign direct investment in Romania in 2019 (BNR, 2019)*

Low values of the net flow of FDI can be observed in the analysed period in relation to the gross domestic product (GDP) in 2011. In 2019, it registered the value of 2.3 percent of GDP, decreasing from 2.6 percent in the last two years, mainly due to the sharp increase in 2019 of nominal GDP expressed in euro (+9.1 percent compared to the previous year).

Multinational companies have an important role in the Romanian economy, they carry out international production activities through FDI. Using transfer pricing, they are concerned with tax optimization. In order for the tax result to be lower and, implicitly, for taxation to be lower, these companies can artificially increase their debts within the group.

According to Baniță (2019), approximately 80% of the companies in the top 100 are multinational companies. Of these, most report low profit compared to turnover, which determines a low profit tax paid by these companies in Romania. Transfer pricing is considered a tool that can help erode the tax base and shift profits.

The evolution of the transfer price adjustments established by ANAF (National Agency for Fiscal Administration) following the fiscal inspections performed for the period 2010-2019 is presented in the following table:

*Evolution of transfer pricing adjustments* Table 3

Year	Additional tax obligations established (Income tax and accessories) - million lei	Reduction of fiscal losses - million lei
2010	13,8	5,9
2011	66,6	188
2012	11	100
2013	4	72
2014	28	60

Year	Additional tax obligations established (Income tax and accessories) - million lei	Reduction of fiscal losses - million lei
2015	192	334
2016	87	428,1
2017	60	316,4
2018*	149,3	237,7
2019	67,2	547,5

Source: ANAF press releases, Annual Performance Reports, [www.anaf.ro](http://www.anaf.ro), (Butnaru and Simionescu, 2018)

\*For 2018, we obtained data related to the period January - November 2018 from the "Budget fiscal strategy for the period 2019-2021" developed by the Romanian Government. December 2018 was estimated quantitatively based on information related to January-November 2018 (The Romanian Government - MFP, 2019).

Regarding the way in which the profits are transferred, in the specialized literature we could find two techniques for moving the profits (Ban and Rusu, 2019), respectively:

➤ *Incorrect transfer prices set between subsidiaries*

This profit shifting technique involves an underestimation of the prices of goods and services in the event of a transfer from related parties located in high tax jurisdictions to related parties located in low tax jurisdictions, or overvaluation in the opposite situation (Hebous and Johannesen, 2015).

➤ *Strategic allocation from the balance sheet*

The strategic allocation technique in the balance sheet involves the transfer of assets that generate profits to related parties located in jurisdictions with low tax rates (Karkinsky and Riedel, 2012), respectively the transfer of liabilities that generate debts to jurisdictions where there are high rates profit taxes (Ruf and Weichenrieder, 2015).

In Romania, we identified studies on direct investments; the relevant ones were those conducted by the Foreign Investors Council (2017 and 2020), such as:

- *Foreign direct investments - their evolution and importance in Romania* (2017). According to this study, foreign direct investment has fundamentally contributed to building a market economy that operates in Romania;
- *Foreign direct investment in Romania* (2020), which analysed the evolution of foreign investment in Romania (value, sectors of activity), making also some proposals in this regard.

### **3. Econometric Model for Testing the Relationship between the Level of Adjustments of Fiscal Obligations related to Transfer Prices and Foreign Direct Investment in Romania**

In this section of the paper, we present the econometric study conducted to analyse the impact of adjustments to tax obligations in the field of transfer pricing on foreign direct investment for the period 2011-2019.

The relationship between investment and GDP has long been studied by economists. We recall the Keynesian theory, according to which investment is the engine of economic growth through the investment multiplier. Considering a closed economy without a government:

$$Y = C + I \quad (1),$$

where C is global consumption, I is global investment and Y is GDP or global income.

$$\Delta Y = \Delta C + \Delta I \quad (2)$$

In Keynes's view, consumption is a function of income, so the change in consumption is equal to the change in income multiplied by the marginal propensity to consume (MPC).

$$MPC = \Delta C / \Delta Y \quad (3)$$

$$\Delta Y = MPC * \Delta Y + \Delta I \quad (4)$$

$$\Delta Y / \Delta I = 1 / (1 - MPC) \quad (5)$$

$$k = 1 / (1 - MPC) \quad (6),$$

where k is the investment multiplier.

Thus, investments are considered the engine of economic growth through their multiplier effect. Obviously, both domestic and foreign direct investment create jobs and improve know-how.

The variables included in our study are the established additional tax liabilities (income tax and accessories), taken largely from the ANAF Annual Performance Reports and adjusted with the GDP deflator and foreign direct investment (source: BNR), which were also adjusted with the GDP deflator. The GDP deflator, with the help of which we adjusted the time series, was taken from the Eurostat database, seasonally and calendar adjusted series.

To maintain the variables stationary, we made logarithms of the time series and calculated the first difference. A stationary time series involves constant averaging and variation over time. To test the stationarity of time series, we used the Augmented Dickey-Fuller test. The probability associated with the stationarity tests is below the chosen significance level of 5% for tax liabilities and below the significance level of 10% and close to the significance level of 5% for real FDI (see Figures 1 and 2 below). Thus, the null hypotheses according to which the tested time series are not stationary are rejected.

Null Hypothesis: DL\_FISCAL\_OBLIGATIONS has a unit root

Exogenous: Constant

Lag Length: 1 (Automatic - based on SIC, maxlag=1)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-4.536915	0.0134
Test critical values:		
1% level	-4.803492	
5% level	-3.403313	
10% level	-2.841819	

\*Mackinnon (1996) one-sided p-values.

Source: Calculations made by the authors

Fig. 1. Testing the stationarity of real tax obligations

Null Hypothesis: DL\_REAL\_FOREIGN\_INVEST has a unit root

Exogenous: Constant

Lag Length: 1 (Automatic - based on SIC, maxlag=1)

	t-Statistic	Prob.*
Augmented Dickey-Fuller test statistic	-3.272626	0.0533
Test critical values:		
1% level	-4.582648	
5% level	-3.320969	
10% level	-2.801384	

\*Mackinnon (1996) one-sided p-values.

Source: Calculations made by the authors

Fig.2. Testing the real FDI stationarity

We analyse the relationship between real foreign direct investment and fiscal obligations by a simple linear regression. The simple linear regression equation has the following form:

$$Y = \alpha + \beta \cdot X + u \quad (7)$$

Null hypothesis:

$$\beta = 0 \quad (8)$$

$$DL\_REAL\_FDI = 0.0361514545142 - 0.014585892713 * DL\_FISCAL\_OBLIGATIONS \quad (9)$$

Dependent Variable: DL\_REAL\_FOREIGN\_INVEST  
Method: Least Squares  
Date: 05/20/21 Time: 20:00  
Sample (adjusted): 2011 2019  
Included observations: 9 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.036151	0.007742	4.669523	0.0023
DL_FISCAL_OBLIGATIONS	-0.014586	0.005732	-2.544594	0.0384
R-squared	0.480518	Mean dependent var		0.034142
Adjusted R-squared	0.406306	S.D. dependent var		0.029986
S.E. of regression	0.023105	Akaike info criterion		-4.504422
Sum squared resid	0.003737	Schwarz criterion		-4.460594
Log likelihood	22.26990	Hannan-Quinn criter.		-4.599002
F-statistic	6.474957	Durbin-Watson stat		2.661914
Prob(F-statistic)	0.038406			

Source: Calculations made by the authors

Fig. 3. The link between FDI and fiscal obligations

The probability associated with the F test is 3.84%, so below the chosen significance level of 5%, thus the null hypothesis that  $\beta = 0$  is rejected.

The coefficient of determination  $R^2$  shows that 48% of the variant of the dependent variable is explained by the independent variable, and the fiscal obligations respectively.

The increase by a percentage in the real fiscal obligations leads to the decrease by 0.0145 percent of the real foreign direct investments.

To confirm that the regression results are valid, we will test the autocorrelation between the normality of the residues and the heteroscedasticity of the errors. Among the consequences of autocorrelation is the oversizing of the coefficient of determination, and the non-normality and heteroscedasticity of the residues imply that the estimators of the parameters in the model do not have the property of maximum likelihood (Jula, 2011).

The Durbin Watson test can take values between 0 and 4. A value around 2 indicates that the residues do not auto-correlate at the first lag. When the test value is below 2, it indicates a positive autocorrelation, and when the values are above 2, the test indicates a negative autocorrelation (Johnston and DiNardo, 1997). In this case, the value is 2.66. To test whether there is a negative residual autocorrelation, we calculated it.

$$\rho_k = \frac{\sum_{i=k+1}^n (x_i - x_{medium}) * (x_i - k - x_{medium}) / (n - k)}{\sum_{i=1}^n (x_i - x_{medium}) / n} \quad (10)$$

$\rho_k$  - the k order of the autocorrelation

n - number of observations

x medium - average of residues (Codirlaşu et al., 2010)

Date: 03/02/21 Time: 18:00  
 Sample: 2011 2019  
 Included observations: 9

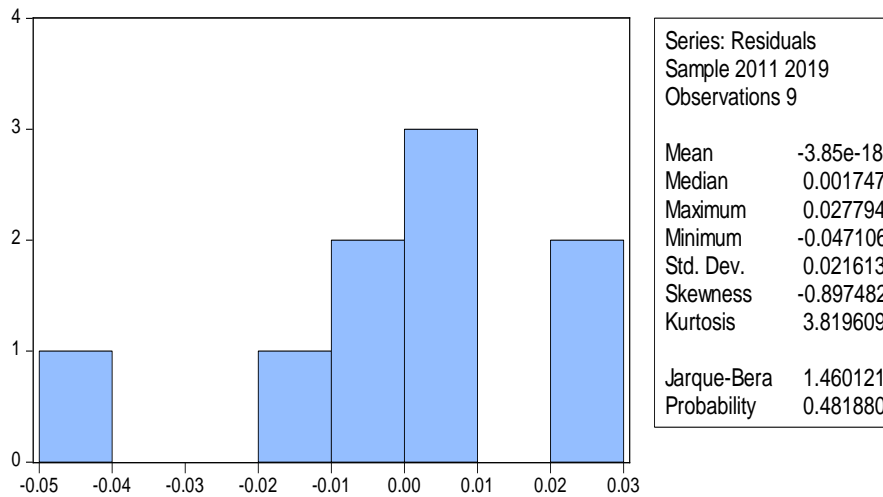
	Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
1	-0.345	-0.345	1.4718	0.225		
2	0.031	-0.100	1.4852	0.476		
3	-0.414	-0.501	4.3160	0.229		
4	0.196	-0.223	5.0761	0.280		
5	-0.044	-0.230	5.1229	0.401		
6	0.143	-0.231	5.7998	0.446		
7	-0.069	-0.174	6.0327	0.536		
8	0.001	-0.220	6.0329	0.644		

Source: Calculations made by the authors

Fig. 4. Residue autocorrelation testing

Given the probabilities obtained in the figure above, we cannot reject the hypothesis that there is no residual autocorrelation.

To test the normality of the residue, we apply the Jarque-Bera test. The test reveals whether the asymmetry coefficients (Skewness) and Kurtotica (Kurtosis) match the values of a normal distribution, respectively 0 for the asymmetry coefficient and 3 for Kurtotica.



Source: Calculations made by the authors

Fig. 5. Residual normality testing

The probability associated with this test is 0.48, above the chosen level of 0.05, which indicates that the residues are normally distributed.

Heteroscedasticity indicates that the residues do not have a constant variance (dispersion). The Breusch-Pagan-Godfrey test regresses the quadratic errors according to a constant and the regressors from the initial equation. The null hypothesis is that the



residues are not heteroscedastic (they are homoscedastic). The probabilities associated with Chi tests are above the chosen significance level of 0.05, so the null hypothesis is not rejected. The residues are homoscedastic.

Heteroskedasticity Test: Breusch-Pagan-Godfrey				
F-statistic	3.192547	Prob. F(1,7)		0.1171
Obs*R-squared	2.819013	Prob. Chi-Square(1)		0.0932
Scaled explained SS	2.404181	Prob. Chi-Square(1)		0.1210
Test Equation:				
Dependent Variable: RESID^2				
Method: Least Squares				
Date: 05/20/21 Time: 20:05				
Sample: 2011 2019				
Included observations: 9				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.000455	0.000220	2.073657	0.0768
DL_FISCAL_OBLIGATIONS	-0.000290	0.000163	-1.786770	0.1171
R-squared	0.313224	Mean dependent var		0.000415
Adjusted R-squared	0.215113	S.D. dependent var		0.000739
S.E. of regression	0.000655	Akaike info criterion		-11.63031
Sum squared resid	3.00E-06	Schwarz criterion		-11.58649
Log likelihood	54.33641	Hannan-Quinn criter.		-11.72489
F-statistic	3.192547	Durbin-Watson stat		1.974912
Prob(F-statistic)	0.117130			

Source: Calculations made by the authors

Fig. 6. Breusch Pagan Godfrey heteroscedasticity test

From a statistical point of view, we proved the existence of a relationship between foreign direct investment and the adjustments of fiscal obligations resulting from the field of transfer pricing.

#### 4. Conclusions and Personal Contributions

From a statistical point of view, through the econometric study we proved the existence of a relationship between foreign direct investment and the adjustments of fiscal obligations resulting from the field of transfer pricing.

We came to the conclusion that an increase of one percent in real tax liabilities leads to a decrease of 0.0145 percent in real foreign direct investment.

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